

Supplement of *Clim. Past*, 10, 1983–2006, 2014  
<http://www.clim-past.net/10/1983/2014/>  
doi:10.5194/cp-10-1983-2014-supplement  
© Author(s) 2014. CC Attribution 3.0 License.



*Supplement of*

## **HadISDH land surface multi-variable humidity and temperature record for climate monitoring**

**K. M. Willett et al.**

*Correspondence to:* K. M. Willett ([kate.willett@metoffice.gov.uk](mailto:kate.willett@metoffice.gov.uk))

## Supplementary Material

### 1) Quality Control of the Hourly Data

The quality control process for HadISDH.2.0.0 is identical to that described for HadISDH.landq.1.0.0 in Willett et al. 2013b (Sect. 2 and 3.1) and fully described in Dunn et al. (2012). This includes three humidity specific tests for supersaturation, wick drying events and issues with static or removed values during extremes of temperature. Of the 3679 stations selected for HadISDH.2.0.0, for  $T_d$ , in total 79.7 % of stations have  $\leq 1$  % of hourly data removed and 98.8 % of stations have  $\leq 5$  % of hourly data removed. For  $T$ , 85.9 % of stations have  $\leq 1$  % of hourly data removed and 98.5 % of stations have  $\leq 5$  % of hourly data removed. Thus  $T_d$  removal percentages exceed those for  $T$ , and this is especially so in the tropics.

### 2) Climatology Selection Criteria

For each station, monthly-hour means are taken for each hour of the day where at least 15 days are present for that hour within the month (e.g., Jan-1976-00hrs, Jan-1976-01hrs etc.). For climatological monthly-hour means there must be at least 15 years of data within the 1976-2005 period for each month of the year and at least 1 year in each decade (1976-1985, 1986-1995, 1996-2005). Climatological monthly means are calculated from climatological monthly-hour means only when there are at least four climatological monthly-hour means across the diurnal cycle, with at least one in each eight hour tercile (00:00–08:00, 08:00–16:00, 16:00–24:00 UTC) of the day.

### 3) Monthly Mean Calculation

Monthly means are calculated from the hourly data by first taking the monthly-hour mean for each hour of the day where there are at least 15 days present within the month. Similarly to the climatology, for a monthly mean to be calculated there must be at least four monthly-hour means per day, with at least one in each eight hour tercile (00:00–08:00, 08:00–16:00, 16:00–24:00 UTC) of the day. This prevents biasing towards night or day, or biases arising from systemically changing observation times aliasing into the record.

### 4) Pairwise Homogenisation Algorithm (PHA) Overview

The PHA (Menne and Williams 2009) detects changepoints by comparing each station with every other station within its neighbour network. These networks contain up to 40 stations with the highest correlations of first differences of climate anomalies with the target station. Changepoints are assigned to a station and time point by identifying statistical breaks in paired series and then deconvolving the matrix of returned changepoints to ascertain the station from which each originates. The neighbours are then used to estimate the size of the identified inhomogeneities. In cases where the adjustment can be identified sufficiently confidently an adjustment is applied relative to the more recent homogeneous subperiod. Results of station

removals during PHA (and ID PHA) processing, and summary adjustment statistics are listed for each variable in Table SM1.

## 5) Application of PHA on $T$ and DPD

First we homogenise  $T$  and DPD directly with PHA. DPD has been chosen over  $T_d$  for initial PHA homogenisation because applying PHA directly to DPD and then calculating  $T_d$  results in fewer cases of unphysical supersaturation than applying PHA to climate anomalies of  $T$  and  $T_d$  (3% of stations versus 5% of stations). There is little difference in the signal-to-noise ratio between DPD and  $T_d$  – changepoints are detected slightly more frequently in DPD than for  $T_d$  (2.53 versus 2.16 changepoints per station). Comparison of applying PHA directly to  $T_d$  as opposed to deriving  $T_d$  from homogenised  $T$ -DPD shows very little difference. While the largest individual gridbox trends occur when  $T_d$  is derived, the directly homogenised regional average trends (significant) are larger (by  $0.01 \text{ } ^\circ\text{C decade}^{-1}$ ) for the globe, Northern Hemisphere and tropics. Over the poorer sampled Southern Hemisphere neither method results in significant trends, varying from  $-0.03 \text{ } ^\circ\text{C decade}^{-1}$  to  $0.01 \text{ } ^\circ\text{C decade}^{-1}$ . Of most interest is the change from positive trends in direct PHA  $T_d$  to negative trends in derived  $T_d$  over coastal southwestern USA. In general, the derived  $T_d$  shows a larger area of significant drying compared to using PHA directly.

Homogenisation statistics for  $T$  and DPD are shown in Table SM1. During direct PHA, 3 and 12 stations are removed from  $T$  and DPD respectively because there are too few highly correlating neighbours to constitute a neighbour network. Further stations are removed for two reasons. Firstly, after direct PHA, which removes subperiods of data when an adjustment size cannot be robustly estimated, some stations no longer have sufficient data to calculate a climatology. Secondly, due to seasonally invariant adjustments there are cases where adjusted values are physically implausible in terms of saturation. This is discussed in Sect. 7.

## 6) Application of ID PHA on $T$ , $T_d$ , $T_w$ , $q$ , $e$ and RH

For each station all changepoints are obtained from the PHA logs for  $T$  and DPD for that station and sorted chronologically. PHA does not allow two changepoints to fall within 12 months of each other and so for ID PHA any changepoint lying within 12 months of an earlier changepoint is ignored. A neighbour network is established for each station from the 40 (maximum, with at least 7) highest correlating neighbours based on the first difference series of the anomalies. Stations with fewer than 7 neighbours are removed at this stage, in addition to the 12 stations removed during direct PHA on  $T$  and DPD, and the 6 stations removed for containing extremely large inhomogeneities (Table SM1). A distribution of potential adjustments is obtained by calculating the difference in medians between the two homogeneous sub-periods surrounding the changepoint for each candidate minus neighbour pair within the network. The median of the distribution becomes the adjustment size and the  $1.65\sigma$  uncertainty is obtained from the 5<sup>th</sup> to 95<sup>th</sup> percentiles, in keeping with direct PHA.

Although  $q$  and  $e$  are derived from hourly  $T_d$  only, changepoints from  $T$  as well as DPD must be used because if there is a simultaneous changepoint with an inhomogeneity of the same size in  $T$  and  $T_d$  then it would not appear in DPD. Also, given historical practices at many synoptic stations in reality the hourly  $T_d$  provided by ISD (and then HadISD) is mostly likely to have originated as  $T_w$  or even RH and then been converted which means that  $T$ , and any inhomogeneities in it, will have been incorporated at some stage in the processing chain. In practice,  $T$  also undergoes a second homogenisation using ID PHA with changepoints from DPD to apply suspected missing changepoints.

Homogenisation statistics are shown in Table SM1. Although no data are removed during ID PHA, a few stations have insufficient monthly means to calculate a climatology (see Sect. 2) and so are now removed. More stations are removed due to saturation problems (Sect. 7).

## **7) Removal of Homogenised Stations with Saturation Problems**

For all variables except  $T$  there is a possibility of homogenisation adjustments resulting in physically unrealistic values in terms of saturation (i.e., RH should not go below 0 %rh or above 100 %rh). For DPD (and  $T_d$  as implicated by DPD), any stations where DPD goes below 0 °C are removed because this implies that  $T_d > T$ . For  $T_w$ , each month is compared with homogenised  $T$  for the same month, and cases where  $T_w > T$  are considered supersaturated. For RH,  $q$  and  $e$ , if the simultaneous homogenised RH value is greater than 100 %rh then all values are considered to be supersaturated. For  $q$ ,  $e$  and RH, a value should never go below zero.

Although this supersaturation is physically possible especially in very cold climates (Makkonen and Laakso, 2005), it is uncommon and in this case very likely an artefact of seasonally invariant adjustments. Any such occurrences lead to removal of the entire station. Removed stations are identified in Figures SM1 and Figure 2 (main text) and counts listed in Table SM1.

There are a very high number of supersaturation occurrences for  $T_w$  resulting in a much smaller number of stations passing through to be part of the gridded product. Users of the  $T_w$  product should be aware of this. It is a minor problem for  $e$ ,  $q$ , and RH where only a few stations are affected and removed. Further work is needed to improve the method in future versions in terms of consistency between  $T$  and  $T_w$ . This could include seasonally varying adjustments and simultaneous multi-variable homogenisation. However, these are significant advancements over current state of the art homogenisation methods.

## **8) Final Station Coverage Prior to Gridding**

After the completion of the climatology requirements, QC and homogenisation, a number of stations are removed, which differs slightly for each variable. Final counts are listed in Table SM1 and final station locations

shown in Figure SM1 and Figure 2 (main text). There are 2494 stations common to all datasets. The loss of supersaturated stations for the  $T_w$  product significantly affects this number, resulting in very few common stations above 60 °N.

The spatial coverage changes over time and this differs by variable. This is shown as missing data and total counts in Figure 3 (main text) and Figure SM2. Total station counts for each WMO region are shown in Figure SM3. It is clear that missing data are a problem for all regions of the world. In some cases these are QC removals but mostly they are data that are missing from the source archives (i.e., ISD). The station selection was initially based on having a minimum data presence across the climatology period (1976-2005). This balances selecting enough stations to create a product that can be considered global versus having reasonably continuous temporal coverage so as not to bias any long-term analysis.

The very steep tail offs for  $T$ , DPD and  $T_d$  (as a result of  $T$  and DPD) are due to the application of direct PHA. The direct PHA removes some data at the start and end of records because it is more difficult to assign a robust adjustment to changepoints located near endpoints. This data removal also leads to a more even distribution of missing months throughout for the direct PHA variables.

There is a clear tail off in station coverage at both the start and ends of the record period for all variables. This is largely due to the initial selection of stations from the ISD database which took place in 2008. There are future plans to revisit this selection process and also the merging process which should improve the tail offs and the 2005 drop off which is clearly due to a loss of USA stations at this time. Figure SM4 shows the locations of USA stations that stop reporting post-2005. While the fall-out of these 128 stations is a concern, especially for users focussing on the USA, the stations are reasonably evenly spread. This reduces the likelihood of any biases being introduced as there are 600+ other stations still reporting.

## 9) A Model for Estimating Uncertainty at the Station Level

The uncertainty estimate for the station provides an individual value for each monthly mean anomaly based on uncertainty in the climatology calculation given missing data (clim), in any homogenisation adjustments made and missed (adj) and also from the initial hourly measurement (ob). All components are described fully in Willett et al. (2013b) and so only the differences necessitated by the multivariable approach will be discussed here. To briefly recap, for any monthly mean anomaly our model is as follows (note Eq. 1 to 7 are in Table 1 of the main text):

$$u_{anom} = \sqrt{u_{clim}^2 + u_{adj}^2 + u_{ob}^2} \quad (1)$$

### 9.1 Inhomogeneity Uncertainty

The inhomogeneity uncertainty is slightly more complex with the addition of ID PHA. In all cases it is a combination of both the uncertainty on any adjustments applied and the uncertainty for any inhomogeneities that likely remain within the data:

$$u_{adj} = \sqrt{u_{applied}^2 + u_{missed}^2} \quad (2)$$

For direct PHA (and ID PHA, retaining consistency with PHA) a  $1.65\sigma$  uncertainty range is provided for each adjustment and this is used here after transformation to  $1\sigma$  standard uncertainty for consistency with the other components. For  $T$  and  $T_d$  the uncertainty estimates arising from the adjustments applied from ID PHA and direct PHA are combined chronologically. The uncertainty in missed adjustments is spatio-temporally static for each variable. It is obtained by assuming that the ‘missing middle’ should be filled (Figure 5, main text). This is done as described in Willett et al. (2013b). A Gaussian curve is fitted to the middle of the actual distribution. Then a best fit distribution is created by selecting the higher of the Gaussian and the actual distribution. In practice, this takes in the wider tails of the actual distribution and the Gaussian middle which infills the ‘missing middle’ of the actual distribution. Differences are then taken between the merged best fit and the actual distribution (blue dotted lines in Figure 5, Main Text). The standard deviation of this difference becomes the missed adjustment uncertainty. Despite ID PHA resulting in a smaller ‘missing middle’, the estimated uncertainty from missed adjustments is similar. For  $T$ ,  $T_w$ ,  $T_d$ ,  $e$ ,  $q$ , RH and DPD these uncertainty estimates are  $0.25\text{ }^\circ\text{C}$ ,  $0.17\text{ }^\circ\text{C}$ ,  $0.26\text{ }^\circ\text{C}$ ,  $0.19\text{ hPa}$ ,  $0.14\text{ g kg}^{-1}$ ,  $1.00\text{ \%rh}$  and  $0.26\text{ }^\circ\text{C}$  respectively.

## 9.2 Measurement uncertainty

Following the BIPM Guide to the Expression of Uncertainty in Measurement (BIPM, 2008), measurement uncertainties are assigned as belonging to one of two categories. Type A evaluation of uncertainties can be estimated from statistical analysis of multiple observations from the same population. Type B evaluation of uncertainties can be estimated from *a priori* knowledge of the measurement apparatus and the measuring conditions. As in Willett et al. (2013b), only Type B uncertainties are used here.

Over a month, type B uncertainties can have randomly varying components,  $u_{rand}$ , whose effect can be reduced by averaging, and components which cause “systematic” errors,  $u_{sys}$  where the effect is not reduced by averaging. We do not provide a specific estimate for  $u_{sys}$  because this is accounted for by homogenisation and  $u_{adj}$  to some extent. Note that this assumption is valid when using the anomalies. However, for the absolute values, our estimated uncertainty is likely to be an underestimate. This is because the most recent homogeneous sub-period may not be truly representative of the ambient humidity or temperature depending on the instrument, shelter and observation quality. Similarly, a station with no changepoints may still be unrepresentative of the true ambient climate while retaining the accurate long-term trend.

To obtain  $u_{rand}$  (the measurement uncertainty estimate in the monthly mean) we have to assume that the point measurement (e.g., hourly) uncertainty is the same for every point measurement within the month. This is because it is not sensible to use the unhomogenised raw hourly values at this point, and the PHA method can only be applied to monthly data. For the point measurement uncertainty estimate  $u_i$  we use a standard uncertainty of 0.15 °C in the wet bulb depression above the ice point (assumed to be 0 °C in this case), extrapolating this below this point where necessary (Willett et al. 2013b), using the assumption that all measurements are taken with an aspirated psychrometer. These are converted to the appropriate variable using Eq. 1 to 7 in Table 1 (main text). The resulting standard uncertainty in RH varies from 1 %rh to 3 %rh, decreasing with increasing  $T$  and increasing with decreasing RH (NPL/IMC, 1996). In addition, a 0.2 °C uncertainty in  $T$  (after Brohan et al. 2006) is used as the estimated  $u_i$  for  $T$ .

For  $T$  and  $T_w$ ,  $u_i$  directly transfers to 0.2 °C and 0.15 °C respectively for all months. For all other variables  $u_i$  will vary depending on the temperature and humidity at each time step. To provide an example, Table SM2 shows the estimated  $u_i$  for all variables at a range of dry bulb temperatures where we assume saturation (for ease of calculation). To obtain  $u_{rand}$ , these values need to be converted to an estimate over the monthly mean by dividing by the square root of the number of measurements made across a month. In the worst-case scenario this should be at least 4 measurements per day within at least 15 days resulting in  $N \geq 60$ :

$$u_{rand} = \frac{u_i}{\sqrt{N}} \quad (3)$$

For conversion of the  $u_i$  in  $T_w$  to a  $u_i$  specific to each of the other humidity variables, homogenised monthly RH and  $T$  are used in addition to the monthly value of that variable. For  $q$  and  $e$ , first the saturation equivalents  $q_s$  and  $e_s$  are calculated for each month by rearranging Eq. (7, main text) with RH set at 100 %rh. Then, using the actual monthly RH a combined RH value plus RH uncertainty  $RH + \Delta RH$  is created by adding a change in RH dependent on the simultaneous monthly  $T$  as shown in Table SM2. We then calculate the monthly value plus uncertainty  $q + \Delta q$  (or  $e + \Delta e$ ) from the  $RH + \Delta RH$  and the saturated value  $q_s$  ( $e_s$ ) using a further rearrangement of Eq. (7, main text). Finally, the original  $q$  (or  $e$ ) value is subtracted from the  $q + \Delta q$  (or  $e + \Delta e$ ) value to obtain the  $u_i$  for that month. For  $T_d$ , the monthly value is converted to  $e$  using Eq. (3, Main Text) and  $e + \Delta e$  is calculated as described above. Then  $T_d + \Delta T_d$  is calculated from  $e + \Delta e$  using a rearrangement of Eq. (3, main text). Subtracting  $T_d$  from  $T_d + \Delta T_d$  provides the  $u_i$  value. In this case we have not used Eq (4, Main Text) (with respect to ice) when  $T$  is less than 0 °C which means that our uncertainty estimate in these cases will be slightly overestimated. For DPD, a simple addition of the  $u_i$  for  $T_d$  for that month and the 0.2 °C  $u_i$  for  $T$  is used. This is a conservative assumption. As the spatio-temporal coverage is not identical across all variables, in cases where there are no simultaneous monthly values for RH and/or  $T$ , values of 80 %rh and 0 °C are used respectively. For all variables,  $u_{rand}$  is estimated from each  $u_i$  using  $N=60$  (Eq. 3).

## **10) Additional Figures and Tables**

Table SM3 lists all fields available for each variable and the minimum and maximum value of that field for the final HadISDH data-product. These data are provided in netCDF format from [www.metoffice.gov.uk/hadobs/hadisdh](http://www.metoffice.gov.uk/hadobs/hadisdh).

Annual climatological averages across the 1976-2005 period are shown in Figure SM5 for wet bulb temperature, vapour pressure and dew point depression. Other variables can be found in Figure 6 (main text). Decadal trends over the 1973-2013 periods are shown in Figure SM6 for wet bulb temperature, vapour pressure and dew point depression. Other variables can be found in Figure 11 (main text).

Table SM1. Homogenisation statistics for all variables. Note that a further 6 stations are removed for all variables due to very large adjustments being found during PHA for  $T$  (714963, 729595, 719410, 026720, 718260, 535880). IH stands for Inhomogeneity.

Variable	station count (start, finish)	No neighbour stations (IDs)	Missing data stations	Subzero error stations	Supersaturation error stations	$\mu$ change-point frequency per station	Absolute IH size $\mu$	Absolute IH size $\sigma$	IH size $\mu$	IH size $\sigma$	5 largest IHs (ID, size)
$T$ (°C)	3679, 3561	3+2 (+9 from DPD) (854690, 910660, 919250, 085010**, 896110**)	98	0	0	3.56 (PHA: 1.34)	0.38 (PHA: 0.73)	0.46 (PHA: 0.57)	-0.02 (PHA: -0.10)	0.60 (PH A: 0.94)	723783 - 4.58, 718440 4.57, 023260 - 4.44, 599480 - 4.05, 766480 - 3.96
DPD (°C)	3679, 3302	9 (689940, 847820, 854880, 889680, 890020, 890220, 916100, 916430, 919430)	213	0	149	2.52	0.98	0.69	-0.01	1.21	556640 7.17, 614970 7.11, 621030 7.05, 442840 6.90, 942380 - 6.34
$T_d$ (°C)	3667*, 3164	2 (085010**, 896110**)	348	0	147	3.62	0.77	0.65	-0.01	1.01	556640 - 7.27, 614970 - 6.76, 726548 6.72,



		6										403400 -
		(085010**,										5.33,
		689060**,										412400 -
<b>e</b>	3667*,	895320**,	17	52	28	3.76	0.41	0.46	-0.01	0.61		5.19,
<b>(hPa)</b>	3558	895710**,										822810 -
		896110**,										4.58,
		911650**),										765770 4.57,
												412680 -
												4.50

\* 12 stations are removed from initial count because they failed PHA for *T* and DPD because they had no neighbour stations.

\*\* Stations removed during indirect PHA because there were fewer than 7 neighbour stations.

Table SM2. Estimates of standard uncertainty in humidity measurements calculated in terms of equivalent psychrometer uncertainty to represent a “worst case scenario”. All uncertainties are based on a 0.15 °C uncertainty in wet bulb depression. Estimates are made for each value assuming saturation at the given temperature and comparing with a value at RH equal to (100 %rh minus the associated uncertainty in RH) for that temperature band. For DPD, the 0.2 C uncertainty in dry bulb temperature is added linearly to the uncertainty in dew point temperature as in a worst case scenario the error in  $T$  and  $T_w$  would oppose. Calculations of specific humidity used equations from Table 1 (main text).

$T$ (°C)	Uncertainty in RH (%rh)	Uncertainty in $q$ (g kg <sup>-1</sup> ) at saturation	Uncertainty in $e$ (hPa) at saturation	Uncertainty in $T_d$ (°C)	Uncertainty in DPD (°C)
-50 and below	15	0.004	0.006	1.309	1.509
-40	15	0.012	0.019	1.428	1.628
-30	15	0.035	0.057	1.553	1.753
-20	10	0.064	0.104	1.094	1.294
-10	5	0.081	0.131	0.577	0.777
0	2.75	0.105	0.169	0.338	0.538
10	1.8	0.138	0.223	0.271	0.471
20	1.35	0.199	0.318	0.219	0.419
30	1.1	0.298	0.470	0.193	0.393
40	0.95	0.459	0.707	0.179	0.379
50+	0.8	0.672	1.000	0.162	0.362

Table SM3 Description of data contained in the HadISDH CF-compliant netCDF file. Gridboxes are 5 ° by 5 ° beginning with centres of -177.5 °W, -87.5 °S. Times count from 1 on January 1973 to 492 on December 2013. Missing data are recorded as -1e30.

Field	Description	Dimension	Maximum and Minimum values						
			$T$ (°C)	$T_w$ (°C)	$T_d$ (°C)	$q$ (g kg <sup>-1</sup> )	$e$ (hPa)	RH (%rh)	DPD (°C)
abs	Monthly mean	72,36,492	-53.18, 39.72	-47.93, 28.51	-54.41, 27.42	0.00, 23.57	0.00, 36.22	4.78, 98.80	0.02, 35.7
anoms	Monthly mean anomaly (seasonal cycle averaged over 1976-2005 climatology removed)	72,36,492	-16.18, 18.21	-16.02, 12.43	-17.37, 16.25	-6.28, 7.16	-10.04, 11.47	-33.62, 38.81	-14.40, 14.80
Std	Standard deviation of all monthly mean anomalies within the gridbox for each month	72,36,492	0.00, 18.10 (100.00)	0.00, 15.72 (100.00)	0.00, 23.92 (100.00)	0.00, 9.36 (100.00)	0.00, 17.27 (100.00)	0.00, 40.72 (100.00)	0.00, 15.50 (100.00)
combinederr	Station uncertainty and sampling uncertainty combined in quadrature to give a 2σ uncertainty	72,36,492	0.07, 4.74	0.07, 4.02	0.08, 5.56	0.04, 3.51	0.06, 5.62	0.33, 27.83	0.07, 5.60

samplingerr	2 $\sigma$ spatial sampling error	72,36,492	0.01, 3.45	0.00, 3.14	0.01, 4.00	0.00, 0.75	0.01, 1.85	0.07, 23.63	0.00, 3.53
rbar	Average inter-site correlation	72,36	0.10, 0.92	0.10, 0.92	0.10, 0.91	0.10, 0.92	0.10, 0.92	0.10, 0.91	0.10, 0.89
sbarSQ	Estimate of mean variance of individual stations in the gridbox	72,36	0.08, 14.79	0.07, 13.43	0.08, 16.77	0.03, 10.00	0.07, 10.00	1.71, 73.90	0.08, 12.55
stationerr	Climatological, adjustment and measurement uncertainty combined in quadrature to give a 2 $\sigma$ uncertainty	72,36,492	0.04, 4.24	0.05, 3.47	0.04, 5.03	0.02, 3.50	0.03, 5.61	0.29, 20.71	0.04, 4.34
adjerr	2 $\sigma$ adjustment uncertainty	72,36,492	0.03, 4.03	0.04, 2.81	0.03, 4.83	0.02, 3.49	0.03, 5.59	0.20, 20.02	0.03, 3.77
obserr	2 $\sigma$ measurement uncertainty	72,36,492	0.00, 0.05	0.00, 0.04	0.00, 1.50	0.00, 0.13	0.00, 0.21	0.02, 3.87	0.01, 1.55
climerr	2 $\sigma$ climatological uncertainty	72,36,492	0.02, 2.45	0.02, 2.04	0.01, 2.85	0.00, 1.15	0.00, 1.83	0.08, 5.43	0.01, 2.83
clims	Monthly climatologies over the 1976-	72,36,12	-46.08, 37.72	-40.76, 27.39	-48.08, 25.91	0.04, 21.88	0.08, 33.66	16.15, 94.48	0.73, 31.08

mean_n_stations	2005 period Total number of stations within the gridbox over the entire record (e.g., contributing to climatology)	72,36	0, 44	0, 36	0, 38	0, 44	0, 44	0, 44	0, 39
actual_n_stations	Actual number of stations in the gridbox for each month	72,36,492	0, 44	0, 36	0, 37	0, 44	0, 44	0, 44	0, 38

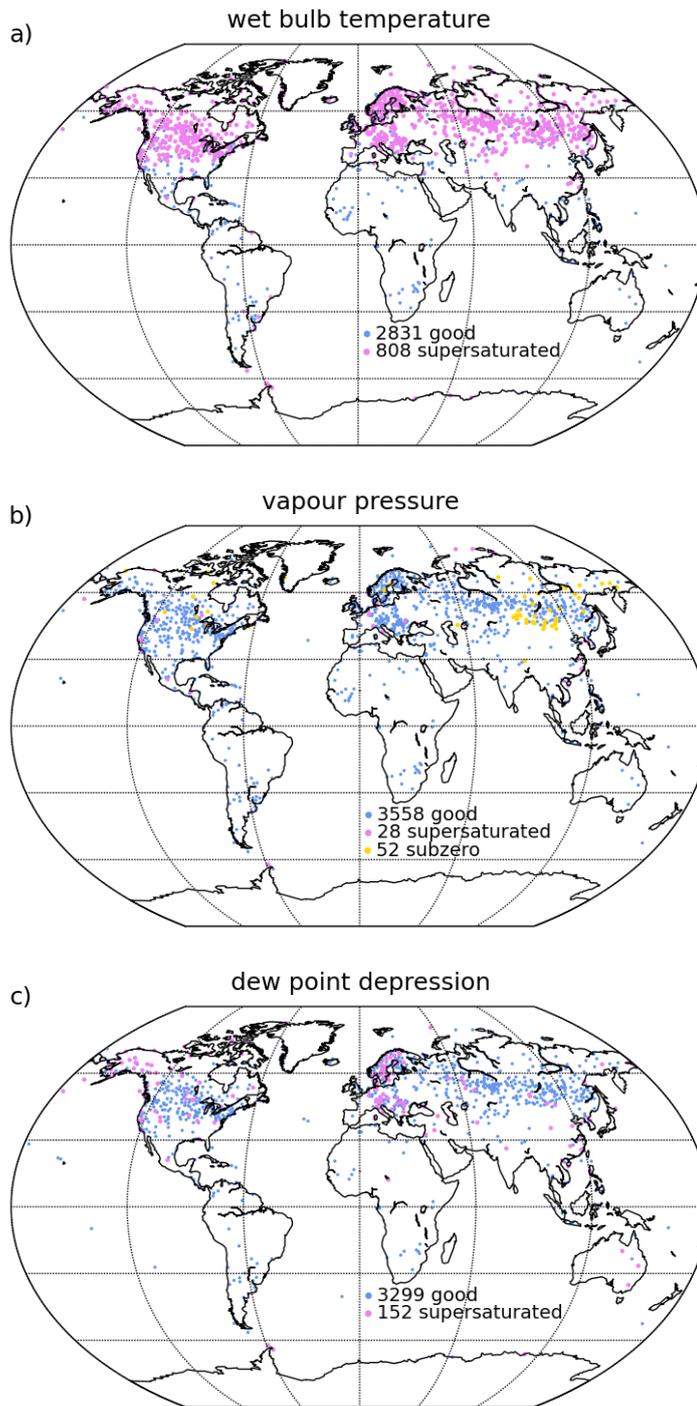


Figure SM1. Additional station coverage maps with supersaturated and subzero stations that have been removed. There are 2494 stations in common to all variables as shown in Figure 2a (main text). a) Wet bulb temperature, b) vapour pressure and c) dew point depression.

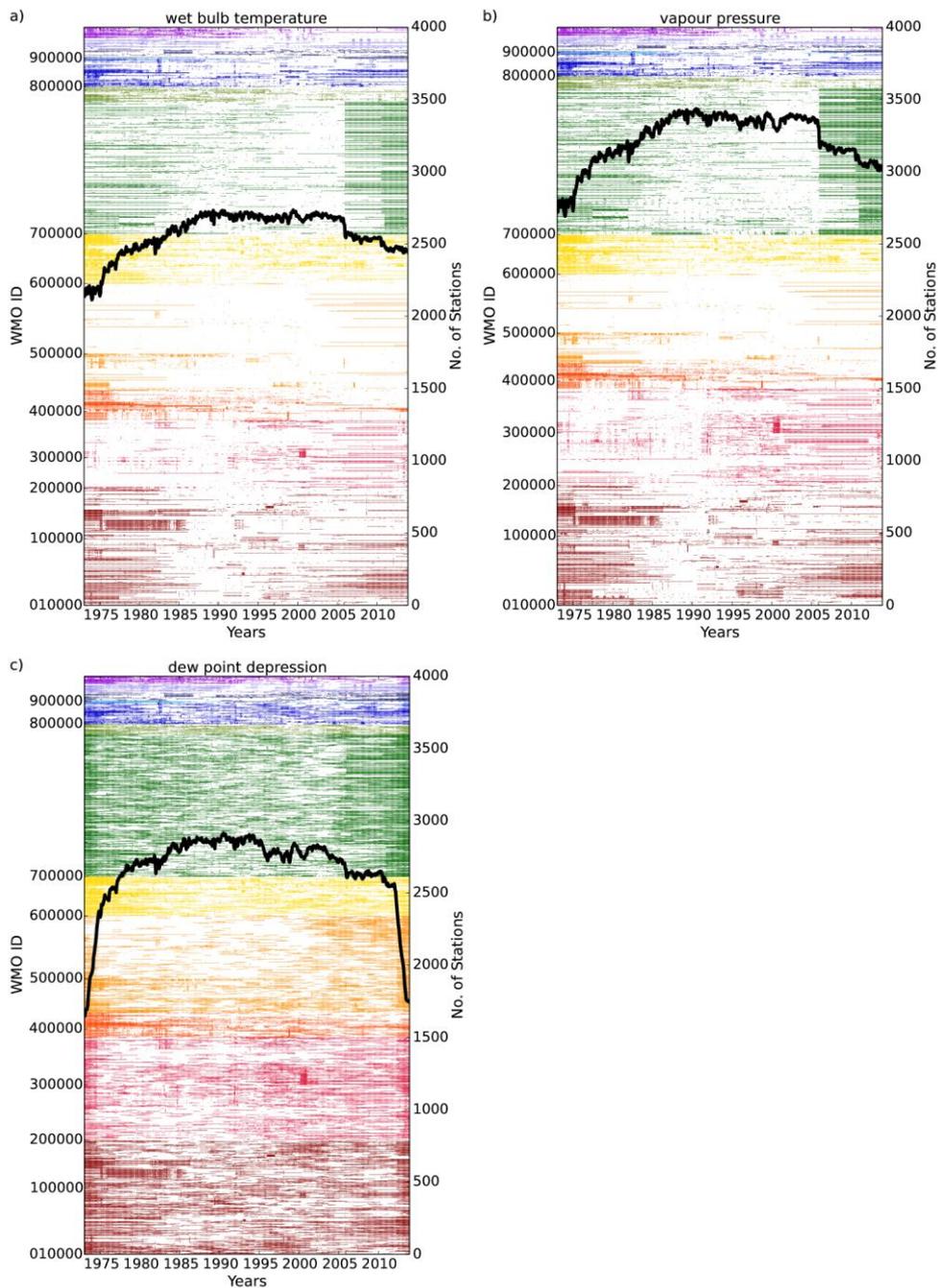


Figure SM2 Station temporal coverage. Colours show missing station months by WMO ID (left y-axis) and the solid black line shows total station count (right y axis). WMO IDs/colours are bounded as follows: Europe: 0-199999 (dark red), Russia/Eastern Europe: 200000-379999 (crimson), Central Asia/Middle East/India/Pakistan: 380000-439999 (dark orange), South East Asia/East Asia/China: 440000-599850 (orange), Africa: 600000-689999 (gold), USA/Canada: 690000-749999 (dark green), Central America: 760000-799999 (olive green), South America: 800000-879999 (royal blue), Antarctica: 880000-899999 (sky blue), Pacific Islands (Inc. Hawaii): 911000-919999 (navy), Papua New Guinea/Australia/New Zealand: 920000-949999 (lilac), Indonesia/Philippines/Borneo: 960000-999999 (purple). a) wet bulb temperature, b) vapour pressure and c) dew point depression. See Figure 3 (main text) for other variables.

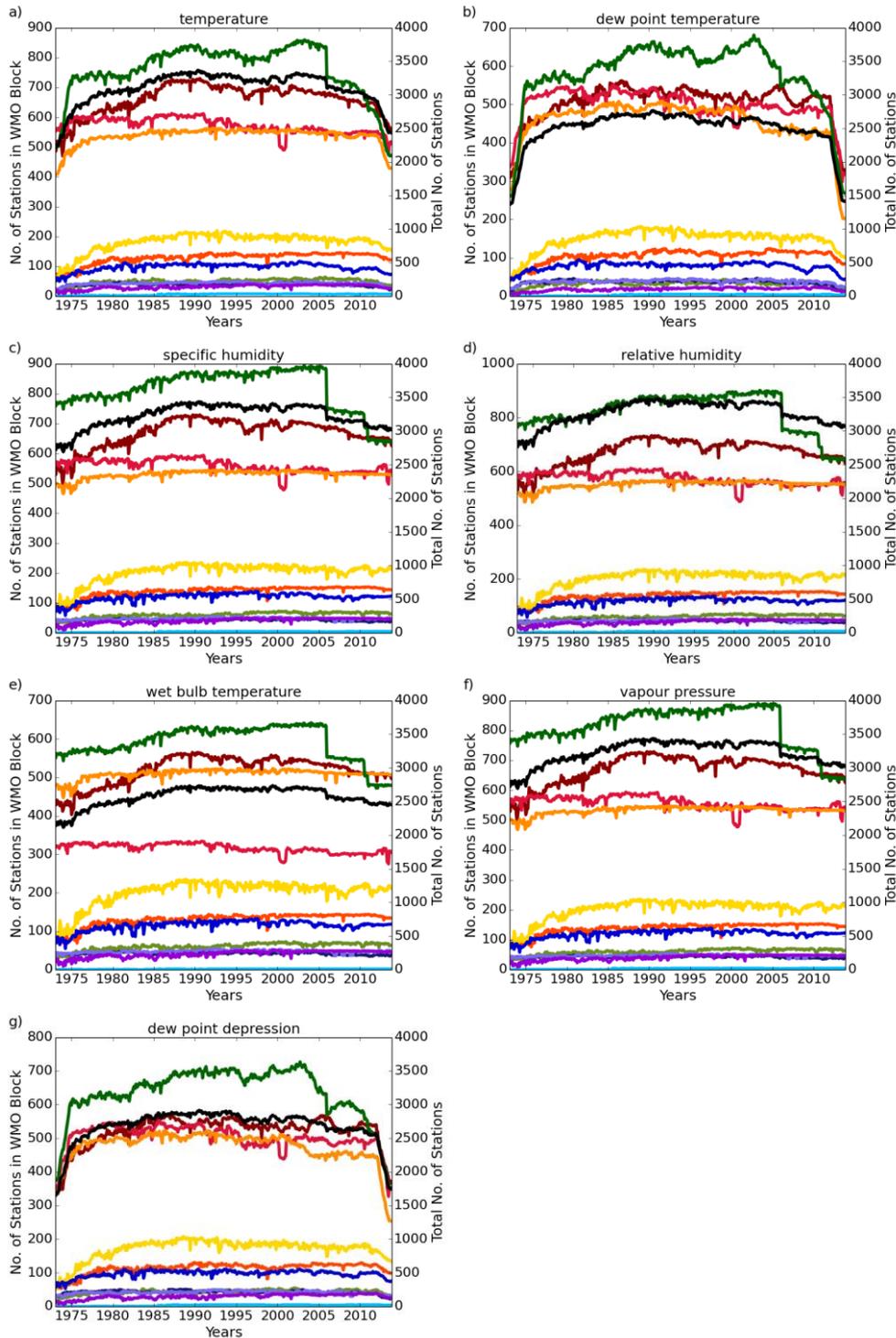


Figure SM3 Station counts per month for each WMO region (colours, left axis) and total for the variable (black line, right axis. WMO regions are bounded as follows: Europe: 0-199999 (dark red), Russia/Eastern Europe: 200000-379999 (crimson), Central Asia/Middle East/India/Pakistan: 380000-439999 (dark orange), South East Asia/East Asia/China: 440000-599850 (orange), Africa: 600000-689999 (gold), USA/Canada: 690000-749999 (dark green), Central America: 760000-799999 (olive green), South America: 800000-879999 (royal blue), Antarctica: 880000-899999 (sky blue), Pacific Islands (Inc. Hawaii): 911000-919999 (navy), Papua New Guinea/Australia/New Zealand: 920000-949999 (lilac), Indonesia/Philippines/Borneo: 960000-999999 (purple). a)

temperature, b) dew point temperature, c) specific humidity, d) relative humidity, e) wet bulb temperature, f) vapour pressure and g) dew point depression.

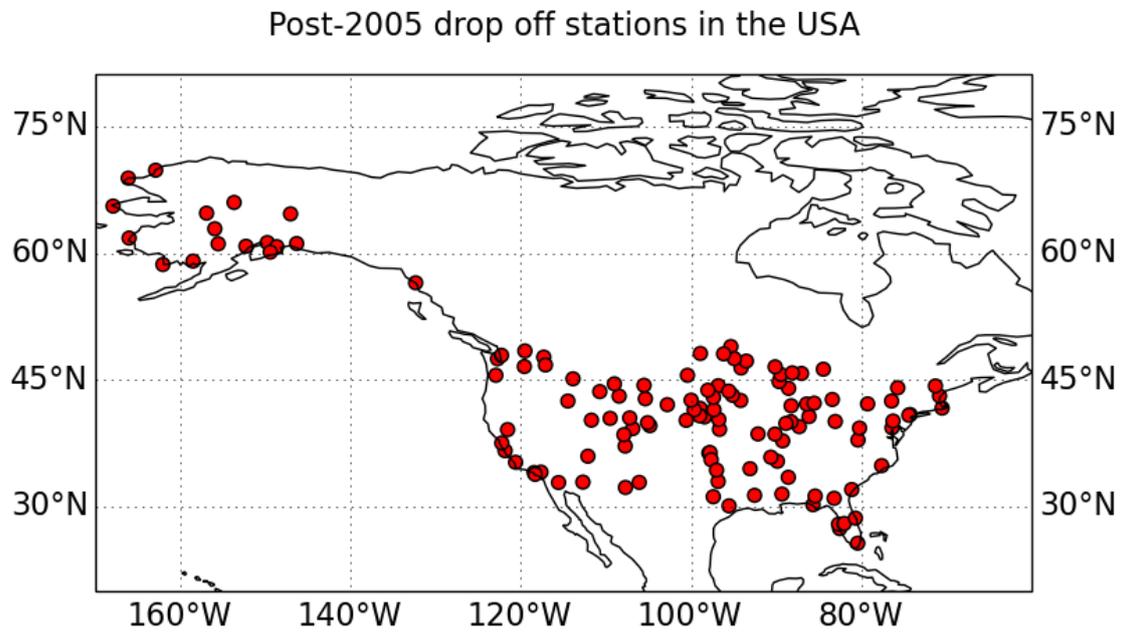


Figure SM4 129 stations in the USA that have no data post-2005.

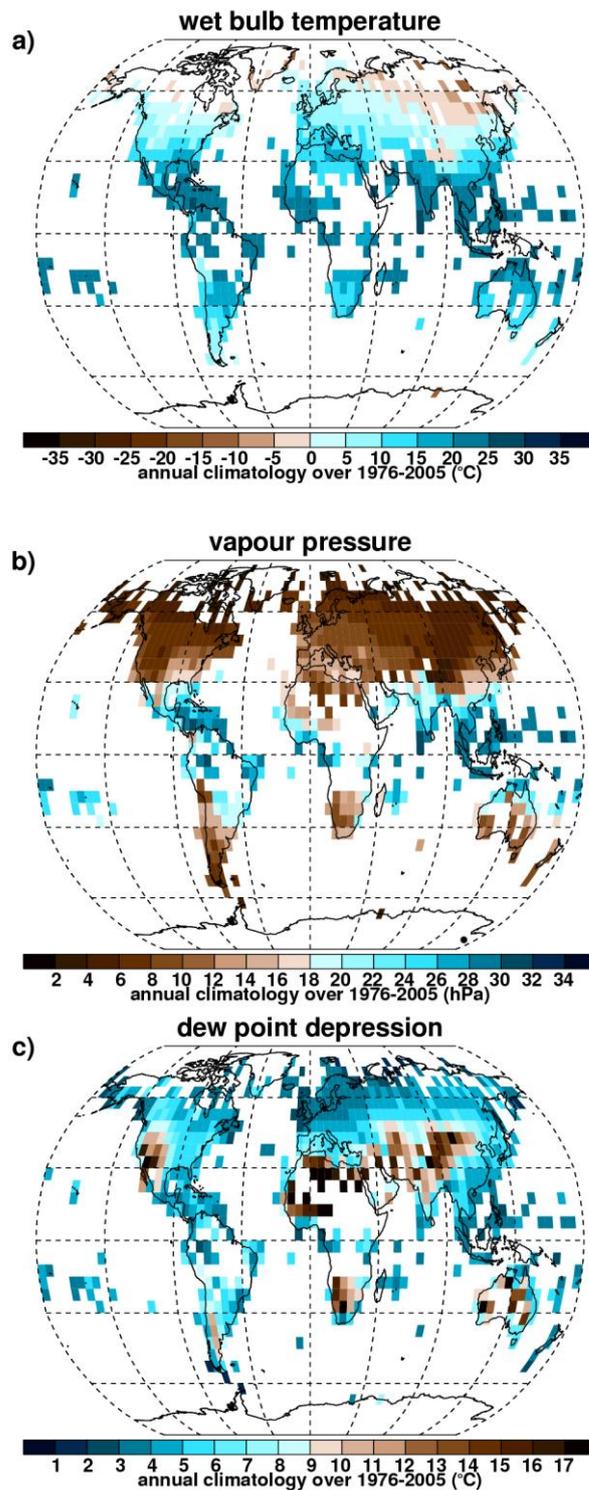


Figure SM5. Annual climatologies (1976-2005) for a) wet bulb temperature, b) vapour pressure and c) dew point depression. See Figure 6 (main text) for other variables.

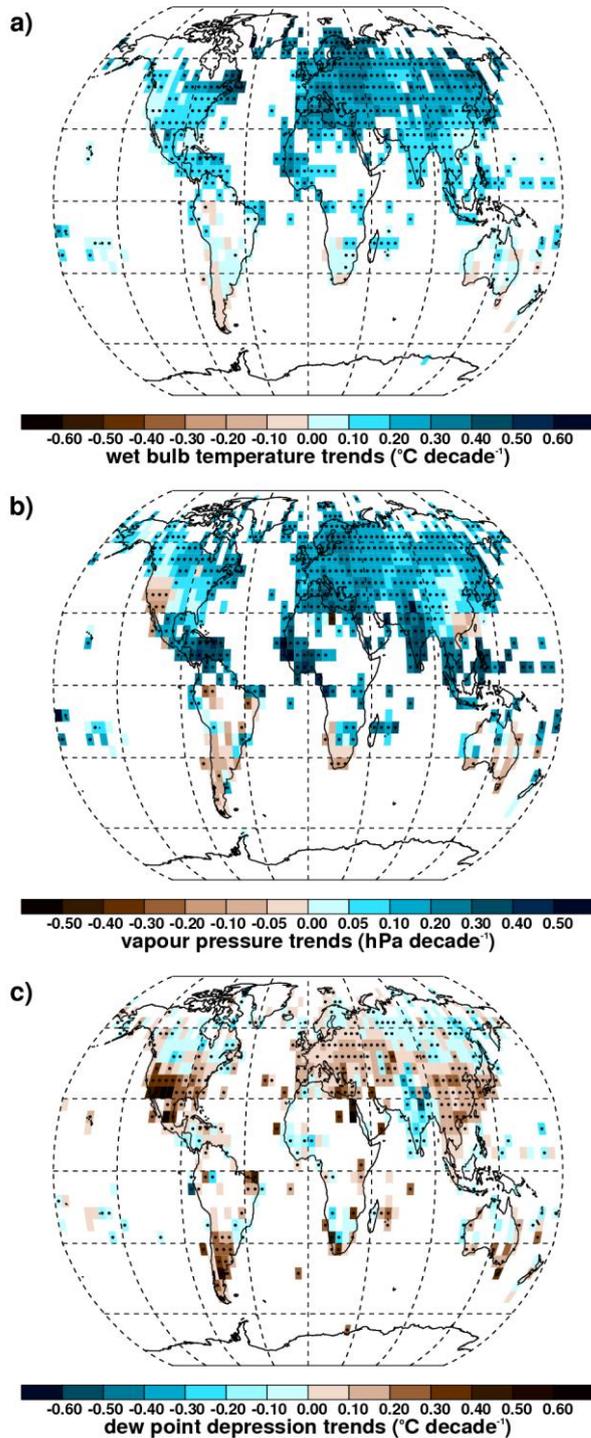


Figure SM6. Decadal trends from 1973 to 2013 for monthly mean climate anomalies relative to 1976-2005. Trends are fitted using the median of pairwise slopes. Black dots show trends that are considered to be significantly different from a zero trend – where the 5<sup>th</sup> and 95<sup>th</sup> percentiles of the pairwise slopes are in the same direction. a) wet bulb temperature, b) vapour pressure and c) dew point depression. See Figure 11 (main text) for other variables.